

## Collision algorithms.

Our next goal is to study a generalization of Shanks' algorithm. We will need to discuss a few problems in probability theory, which are of some interest in their own right.

### The birthday paradox.

- (1) What is the probability that out of a random group of  $n$  people someone has birthday on a given day?
- (2) What is the probability that at least two people share the same birthday?  
(assume the year has 365 days)

# Answers:

(1) It is easier to compute the probability that no one has his/her birthday on the given day (your birthday) and use that the sum of probabilities of complementary events is 1:

$$P(A) + P(A^c) = 1.$$

$P(\text{no one has birthday on the given day}) = \left(\frac{364}{365}\right)^n$  (the birthdays are independent and any day, except the given one, works).  
Hence, we get  $1 - \left(\frac{364}{365}\right)^n$ .

(2) Again, using similar logic,

$$P(\text{two people have the same birthday}) = 1 - P(\text{all } n \text{ people have different birthdays}) =$$
$$= 1 - \frac{365}{365} \cdot \frac{364}{365} \cdot \frac{363}{365} \cdots \frac{365 - (n-1)}{365}$$

↑  
no restrictions on the birthday of the 1<sup>st</sup> person

↑  
the second person can have a birthday on any day except the first person's birthday...

Example. My birthday is July 4<sup>th</sup> (it actually is :)). We have 15 students in the class. What is the probability that someone has birthday on 07/04 as well?

Answer:  $1 - \left(\frac{364}{365}\right)^{15} \approx 0.04$  (4%).

The probability that two of you have birthdays on the same day is

$$1 - \frac{365}{365} \cdot \frac{364}{365} \cdot \dots \cdot \frac{351}{365} \approx 0.25 = 25\%$$

The second answer suggest that the probability of the corresponding event is a lot higher than that of the first one. As the events look similar at a first glance, but their probabilities are very different, such an instance is referred to as paradox.

Thm. An urn contains  $N$  balls (red and blue),  $n$  are red and  $N-n$  are blue. We randomly pick a ball from the urn, record its color and put it back. The procedure is repeated  $m$  times.

(a) The probability that at least one red ball will be picked is  $P(X \geq 1) = 1 - \left(1 - \frac{n}{N}\right)^m$ .

(b) The lower bound is

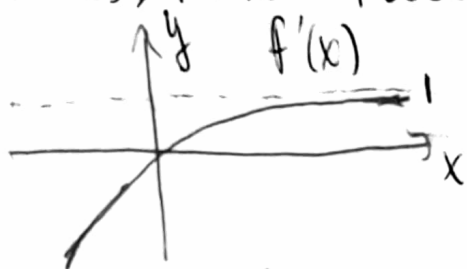
$$P(X \geq 1) \geq 1 - e^{-\frac{mn}{N}}$$

Proof: (a) very similar to the 'birthday paradox' (think it through!).

(b) We first show that  $e^{-x} \geq 1-x$  for  $\forall x \in \mathbb{R}$ .

Consider the function  $f(x) := e^{-x} - 1 + x$ . We need to show that  $f(x) \geq 0$  for all  $x \in \mathbb{R}$ .

Notice that  $f(0) = 1 - 1 + 0 = 0$ . It is sufficient to check that  $f(x)$  decreases for  $x < 0$  and increases for  $x > 0$ . As  $f'(x) = -e^{-x} + 1 \geq 0$  for  $x \geq 0$  and  $f'(x) \leq 0$  for  $x \leq 0$ , the assertion follows.



Using the inequality for  $x = \frac{n}{N}$ , we get

$$e^{-\frac{n}{N}} \geq 1 - \frac{n}{N} \Rightarrow e^{-\frac{nm}{N}} \geq \left(1 - \frac{n}{N}\right)^m \Leftrightarrow -e^{-\frac{nm}{N}} \leq -\left(1 - \frac{n}{N}\right)^m$$

$$\Leftrightarrow 1 - \left(1 - \frac{n}{N}\right)^m \geq 1 - e^{-\frac{nm}{N}} \quad x^m \text{ is monotonically increasing}$$

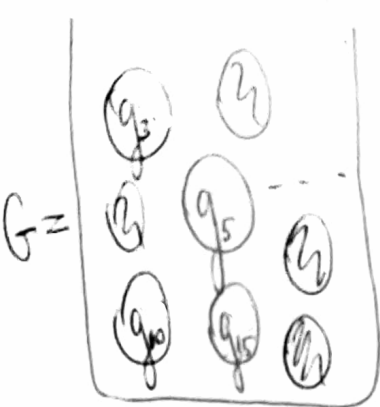
□

Application: let  $G$  be a finite group and  $L_1, L_2 \subset G$  two subsets (not necessarily subgroups) of cardinality  $n < N$ . Then the probability that the intersection of  $L_1$  and  $L_2$  is nonempty satisfies the inequality

$$P(L_1 \cap L_2 \neq \emptyset) \geq 1 - e^{-\frac{n^2}{N}}$$

(we assume that the el-ts of  $L_1$  and  $L_2$  are chosen randomly).





The probability that the intersection of  $L_1$  and  $L_2$  is nonempty is the same as probability that one of elements in  $L_2$  is in  $L_1$  as well. The latter is equal to the probability

of drawing at least one red ball (el-t of  $L_1$ ) in  $n$  (cardinality of  $L_2$ ) tries. Hence, it satisfies the inequality

$$P(L_1 \cap L_2 \neq \emptyset) \geq 1 - e^{-\frac{n^2}{N}}$$

For instance, if  $n = \sqrt{N}$ , then  $P(L_1 \cap L_2 \neq \emptyset) \geq 1 - e^{-1} \approx 63.2\%$  while if  $n = \sqrt{3N}$ , then  $P(L_1 \cap L_2 \neq \emptyset) \geq 1 - e^{-3} \approx 99.9877\%$ .

Remark. Unlike the setup of Shanks' algorithm, we are not guaranteed that the lists

$$L_1 = \{g_1, g_2, \dots, g_n\} \text{ and } L_2 = \{s_1, s_2, \dots, s_n\}$$

(for an arbitrary group  $G$ ) have a nonempty intersection.

### Collision algorithm for DLP.

Proposition. Let  $G$  be a group and  $g$  an element of order  $N$ . Assume we have a well-defined PLP, i.e.  $g^s = h$  (with  $h$  being given). Then the solution ( $s$ ) can be found in  $O(\sqrt{N})$  steps (each step is exponentiation in  $G$ ), using the collision algorithm.

Proof. We mimic the approach in Shank's algorithm.

Write  $s \equiv k - m$  and look for a solution of DLP in the form

$$g^k = h \cdot g^m$$

We make a random choice of  $n$  numbers  $\{a_1, \dots, a_n\}$  and set the first list to be  $L_1 = \{g^{a_1}, g^{a_2}, \dots, g^{a_n}\}$ . Next, randomly pick another  $n$ -tuple  $\{b_1, b_2, \dots, b_n\}$  and set

$$L_2 = \{h \cdot g^{b_1}, h \cdot g^{b_2}, \dots, h \cdot g^{b_n}\}.$$

The 'urn'  $S$  will be the set  $\{1, g, g^2, \dots, g^{N-1}\}$ .

Notice that  $h = g^s$  implies  $L_1, L_2 \subset S$  ( $L_1$  and  $L_2$  are subsets of  $S$ ).

② Once again, no one can guarantee  $L_1 \cap L_2 \neq \emptyset$ , but for  $n$  sufficiently large (say,  $n \sim 2.5\sqrt{N}$  or  $3\sqrt{N}$ ) the probability  $P(L_1 \cap L_2 \neq \emptyset)$  is very close to 1.

Let  $x \in L_1 \cap L_2$ , then  $x = g^k = h \cdot g^m$  (for some  $k \in \{a_1, a_2, \dots, a_n\}$  and some  $m \in \{b_1, \dots, b_n\}$ ).

Then  $s \equiv k - m \pmod{N}$  is the solution to DLP.

## Pollard's $f$ method.

Goal: using the collision type algorithms requires storing a lot of numbers ( $\sim 2\sqrt{N}$  as  $|L_1| = |L_2| \approx \sqrt{N}$ ), which becomes an issue for large values of  $N$ . The algorithm invented by Pollard allows to bypass this issue.

Let  $S$  be a finite set and  $f: S \rightarrow S$  a function. Suppose we choose a point  $x = x_0 \in S$  and iteratively apply  $f$  to it, thus getting the sequence

$$x_0 = x$$

$$x_1 = f(x)$$

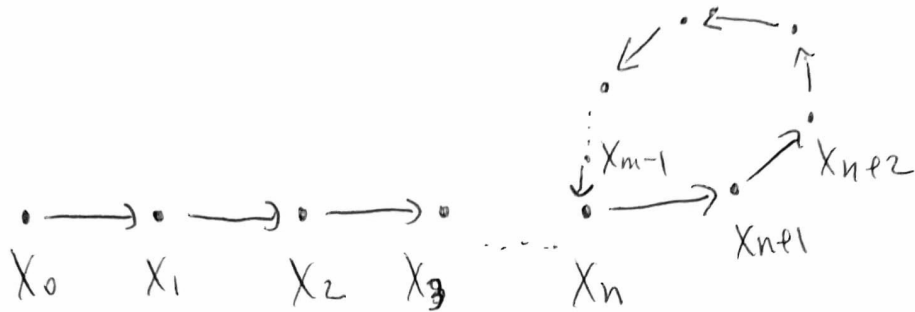
$$x_2 = f(x_1) = f \circ f(x)$$

$$\vdots$$
$$x_n = \underbrace{f \circ f \circ \dots \circ f}_n(x) = f^{(n)}(x)$$
$$\vdots$$

of points in  $S$ .

Rmk. The pair  $(S, f)$  is called a discrete dynamical system.

Rmk. As  $S$  is finite, there must exist a pair of indices  $(m, n)$ ,  $m > n$ , such that  $x_m = x_n$ .



Notation:  $T := \min_n \{n \mid \exists m: x_m = x_n\}$  (tail)

$L :=$  length of the loop  $= \min_n \{n \mid x_{T+n} = x_T\}$ .

Rmk. The shape of the picture above resembles  $\beta$ , hence, the name of the method.

We introduce one more sequence:

$$y_0 = x_0$$

$$y_1 = f \circ f(x) = x_2$$

$$y_2 = f \circ f \circ f \circ f(x) = x_4$$

$\vdots$

$$y_n = x_{2n}$$

Obvious (but important) lemma.

$$x_j = x_i \Leftrightarrow \begin{cases} i \geq T \\ j \equiv i \pmod{L} \end{cases}$$

Corollary.  $y_i = x_i \Leftrightarrow \begin{cases} i \geq T \\ i \equiv 0 \pmod{L} \\ (j = 2i - i = i). \end{cases}$

Thm (Pollard's  $\rho$  Method). Let  $S$  be a finite set,  $|S|=N$ ,  $f: S \rightarrow S$  a map,  $x \in S$  a chosen 'starting element'.

(a) If the sequence  $\{x_0, x_1, x_2, \dots\}$  has a tail of length  $T$  and loop of length  $L$ , then

$$x_{2i} = x_i \text{ for some } 1 \leq i \leq T+L.$$

(b) If the map  $f$  is sufficiently random, then the mean value<sup>(x)</sup> of min.  $i$ , satisfying (a) is

$$E(i) \approx 1.25 \sqrt{N}.$$

(x): among different choices of the starting point.

Rmk. At each moment in time we only need to store  $x_i$  and  $y_i = x_{2i}$ .

Example: DLP for  $\mathbb{F}_p^*$ .

Let's see how to make Pollard's algorithm work in case  $S = \mathbb{F}_p^*$  and how it helps for solving (reducing the storage) the DLP:  $g^s \equiv h \pmod{p}$

Step 1. We choose the function

$$f(x) := \begin{cases} g \cdot x, & 0 \leq x < p/3 \\ x^2, & p/3 \leq x < 2p/3 \\ hx, & 2p/3 \leq x < p \end{cases}$$

Rmk. It is not known if  $f(x)$  is sufficiently random!  
But experimentally ok.

We start with the element  $x_0 = x = 1$ , so

$$x_n = f^{on}(x) = g^{\alpha_n} \cdot h^{\beta_n} \text{ with}$$

$$\alpha_{n+1} = \begin{cases} \alpha_{n+1} \\ 2\alpha_n \\ \alpha_n \end{cases}, \quad \beta_{n+1} = \begin{cases} \beta_n & 0 \leq x_n < p/3 \\ 2\beta_n & p/3 \leq x_n < 2p/3 \\ \beta_{n+1} & 2p/3 \leq x_n \leq p \end{cases}$$

Similarly,  $y_0 = x_0 = 1$ ,  $y_n = g^{\gamma_n} \cdot h^{\delta_n}$  and the collision

$$y_n = x_{2n} \text{ implies } g^{2\alpha_n} \cdot h^{\beta_n} = g^{\gamma_n} \cdot h^{\delta_n} \text{ or}$$

$$(*) \quad g^u \equiv h^v \pmod{p}, \text{ where } u \equiv 2\alpha_n - \gamma_n \text{ and } v \equiv \delta_n - \beta_n \pmod{p-1}.$$

The congruence (\*) implies  $u \equiv v \log_g h \pmod{p-1}$ .

In case  $\gcd(v, p-1) = 1$ , we have that  $v$  is invertible modulo  $p-1$  and  $\log_g h \equiv u \cdot v^{-1} \pmod{p-1}$ .

The case  $\gcd(v, p-1) \geq 2$  requires a bit more work (see page 241 in the book). Also see page 242 for a concrete example (with actual numbers).